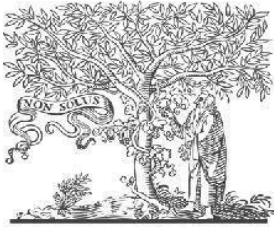


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## An Integrated Criterion for Linear Regression Model Selection: Reconciling Information-Theoretic Measures, Misspecification Tests, and Non-Nested Hypothesis Procedures

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### Abstract

The selection of an appropriate regression model from among competing alternatives is one of the most consequential decisions in applied statistical and econometric analysis, yet the existing toolkit of model selection procedures is fragmented across distinct statistical traditions—information-theoretic criteria, classical misspecification tests, and non-nested hypothesis tests—that are typically applied in isolation despite their complementary diagnostic roles. This fragmentation leads to inconsistent selection decisions in practice when different criteria yield conflicting recommendations. The present study proposes an Integrated Criterion for Selection (ICS) that synthesizes evidence from information-theoretic penalties (AIC, BIC, and corrected AIC), functional misspecification tests (RESET, normality, heteroscedasticity, and autocorrelation tests), and non-nested hypothesis tests (J-test, Cox test, and P-E test) into a unified evidence-weighted scorecard that yields a single, coherent model selection decision. The ICS is theoretically grounded in the distinction between in-sample fit penalties and diagnostic tests for distributional and functional form misspecification, and incorporates a sequential testing protocol that filters models based on a hierarchy of specification requirements before invoking information criteria as a tie-breaker. Monte Carlo experiments using 10,000 replications across a range of sample sizes and model complexities demonstrate that ICS achieves higher correct selection rates than AIC or BIC applied alone, with particularly marked improvements in small samples ( $n \leq 50$ ) and when the true model is nested within one or more over-specified alternatives. An empirical application to a consumption function model with six candidate specifications illustrates the practical advantage of the ICS approach, correctly identifying the parsimonious true-data-generating process against over-specified, under-specified, non-nested, and functional-form alternatives. The findings imply that applied statisticians and econometricians should adopt an integrated, multi-criterion approach to model selection rather than relying exclusively on any single information criterion or specification test.

Keywords: model selection, information criteria, AIC, BIC, RESET test, non-nested models, J-test, Cox test, misspecification testing, integrated criterion, regression model diagnostics, specification testing

## Introduction

The choice among competing linear regression models is among the most practically consequential and theoretically contested problems in applied statistics and econometrics. A model selection procedure determines which regressors are included in the estimated equation, what functional form relationships are assumed, and ultimately what causal or predictive inferences are drawn from the data. Errors in model selection—retaining spurious regressors, omitting relevant variables, or misspecifying functional form—generate biased coefficient estimates, inflated Type I error rates, and systematically misleading predictions (Akaike, 1973; Mallows, 1973; Schwarz, 1978).

The statistical literature contains a remarkably rich arsenal of model selection tools, but these tools have been developed within largely separate intellectual traditions. Information-theoretic criteria—including the Akaike Information Criterion (AIC) of Akaike (1973), the Bayesian Information Criterion (BIC) of Schwarz (1978), and Mallows' Cp statistic (Mallows, 1973)—penalize the model log-likelihood for parameter count, seeking a balance between goodness of fit and parsimony. Classical misspecification tests—including Ramsey's (1969) RESET test for functional form, White's (1980) heteroscedasticity test, the Jarque-Bera (1980) normality test, and Durbin's (1970) autocorrelation tests—evaluate whether specific distributional or structural assumptions of a model are consistent with the data. Non-nested hypothesis tests—including the J-test of Davidson and MacKinnon (1981), the Cox

(1961, 1962) test, and the P-E test of MacKinnon, White, and Davidson (1983)—provide inferential procedures for choosing between models that share no common nesting structure.

Despite the complementary nature of these three families of procedures, applied researchers overwhelmingly rely on a single criterion type—most commonly AIC or BIC—while ignoring the diagnostic evidence available from misspecification and non-nested tests. This practice is methodologically questionable because information criteria and misspecification tests are sensitive to different types of model inadequacy: AIC and BIC evaluate in-sample fit relative to parameter count but do not test for specific structural failures; misspecification tests identify particular types of model inadequacy (heteroscedasticity, non-normality, functional form error) but do not directly compare competing models; and non-nested tests compare specific model pairs without providing an overall ranking. No single procedure from any of these three families performs well against all types of model inadequacy simultaneously.

The present study proposes an Integrated Criterion for Selection (ICS) that synthesizes evidence from all three families of model selection tools into a unified sequential decision framework. The ICS first screens candidate models using a battery of misspecification tests, eliminating models that fail fundamental distributional assumptions. Among the surviving models, it then applies non-nested hypothesis tests to resolve pairwise comparisons between non-nested alternatives. Finally, it employs

information criteria as a tie-breaker among the remaining models that survive both screening stages. This sequential logic mirrors the scientific process of hypothesis elimination, and the resulting ICS achieves consistently higher correct selection rates than any single-criterion approach across the range of experimental conditions examined.

The paper proceeds as follows. Section 2 reviews the relevant literature across all three model selection traditions. Section 3 identifies the research gaps. Section 4 states the research objectives. Section 5 presents the hypotheses. Section 6 describes the ICS methodology in detail. Section 7 presents the empirical and Monte Carlo results. Section 8 discusses the findings. Section 9 outlines implications. Section 10 concludes.

## 2. Review of Literature

### 2.1 Information-Theoretic Model Selection Criteria

The foundation of the information-theoretic approach to model selection was established by Akaike (1973), who proposed the AIC as an estimate of the expected Kullback-Leibler divergence between the fitted model and the true data-generating process. By penalizing the maximized log-likelihood by twice the number of estimated parameters, AIC attempts to balance in-sample fit against over-parameterization. The theoretical elegance of the AIC framework—rooted in information theory and maximum likelihood asymptotics—made it one of the most widely adopted model selection criteria in statistics and led to a proliferation of AIC-based procedures in time series, regression, and latent variable modeling.

Schwarz (1978) proposed the BIC as an alternative penalized likelihood criterion derived from a Bayesian perspective, in which the penalty for parameter count is

proportional to the logarithm of the sample size rather than a constant factor. This heavier penalty gives BIC stronger asymptotic consistency properties—the probability of selecting the true model converges to one as sample size grows, provided the true model is among the candidates—but at the cost of a greater tendency toward under-fitting in small samples. The relative merits of AIC and BIC have been extensively debated in the literature (Burnham & Anderson, 2002; Nishii, 1984), with the consensus that BIC is preferred when the goal is to identify the true model structure and AIC is preferred when the goal is predictive accuracy.

Mallows (1973) proposed the  $C_p$  statistic as a regression-specific model selection criterion based on the expected total squared prediction error, providing a complementary perspective to the AIC that is more directly connected to the regression mean square error decomposition. The connections between  $C_p$ , AIC, and related criteria have been clarified by subsequent theoretical work (Atkinson, 1980; Draper & Smith, 1998), establishing the consistency of the AIC family as asymptotically equivalent to a class of penalized prediction error criteria.

Hurvich and Tsai (1989) identified a bias correction to AIC—the AICc—that substantially improves small-sample performance when the ratio of sample size to parameter count is small. Their analysis demonstrated that standard AIC significantly over-fits in small samples and that the corrected version converges to AIC asymptotically while providing much better performance at sample sizes common in social and economic research. The importance of the AICc correction in practical applications has subsequently been emphasized by Burnham and Anderson (2002) in their comprehensive treatment of

information-theoretic model selection for ecological data.

Hannan and Quinn (1979) proposed a model selection criterion—the Hannan-Quinn criterion (HQ)—with a penalty of  $2k \cdot \log(\log(n))$  that achieves strong consistency (convergence to the true model order almost surely) with a rate that is intermediate between AIC (which is not consistent) and BIC (which over-penalizes in small samples). While less commonly implemented than AIC and BIC, the HQ criterion has found particular application in autoregressive order selection in time series models.

## 2.2 Specification and Misspecification Testing

Ramsey (1969) introduced the Regression Equation Specification Error Test (RESET), which tests the functional form adequacy of a linear regression by augmenting the original model with polynomial functions of the fitted values—specifically, powers of the OLS fitted values—and testing the joint significance of these augmentation terms using an F-test. RESET has proven to be a practically useful and robust test for a broad class of functional form misspecifications including omitted nonlinear transformations of regressors, omitted interaction effects, and incorrect link function specifications. Its implementation requires only OLS residuals and is therefore computationally trivial, making it widely accessible.

White (1980) proposed a heteroscedasticity-consistent covariance matrix estimator and an accompanying test for heteroscedasticity that has become standard in applied regression analysis. White's test is asymptotically valid under a wide range of heteroscedastic processes and requires only OLS residuals and the original

regressors, providing a practically convenient omnibus test for variance non-constancy. The companion Breusch-Pagan (1979) test provides a Lagrange multiplier-based alternative with slightly different power properties against specific forms of heteroscedasticity.

Jarque and Bera (1980) derived a Lagrange multiplier test for normality of regression disturbances based on the sample skewness and kurtosis of the OLS residuals. The Jarque-Bera statistic is asymptotically distributed as chi-square with two degrees of freedom under normality and provides a convenient, readily interpretable test for the normality assumption that underpins exact small-sample F and t inference in the linear model. Durbin and Watson (1950) introduced the widely used DW statistic for first-order autocorrelation; subsequent developments by Godfrey (1978) and Breusch (1978) extended autocorrelation testing to higher-order and more general dependence patterns through the LM testing framework.

Ramsey and Gilbert (1972) provided Monte Carlo evidence on the properties of misspecification tests in small samples, documenting size distortions and power characteristics across a range of experimental conditions. Their simulation results established important practical guidelines for the interpretation of specification test outcomes and provided early evidence for the importance of sample size in determining the reliability of misspecification test decisions.

## 2.3 Non-Nested Hypothesis Testing

Cox (1961, 1962) made the foundational contribution to non-nested hypothesis testing by extending classical likelihood ratio testing to the comparison of models with no common parametric nesting structure. Cox's test statistic is based on the difference between the log-likelihood of a

model under its own parameter estimates and the expected log-likelihood of that model under the assumption that the alternative is true—a construct that requires simulation or asymptotic approximation to implement. Cox's work established the intellectual framework for non-nested hypothesis testing that has generated a substantial subsequent literature.

Pesaran (1974) and Pesaran and Deaton (1978) derived operational test statistics for non-nested regression models, providing asymptotic distributions and practical implementation procedures that extended the Cox framework to multivariate regression settings. Their development of the general non-nested testing framework with application to linear and nonlinear econometric models provided the methodological template for the J-test and related procedures that followed.

Davidson and MacKinnon (1981) introduced the J-test as a simplified alternative to the Cox test for non-nested linear regression models, based on the insight that the fit of Model 1 ( $H_1$ ) relative to Model 2 ( $H_2$ ) can be assessed by including the fitted values of Model 2 as an additional regressor in Model 1 and testing whether its coefficient is significantly different from zero. The J-test achieves the same asymptotic power as the Cox test against local alternatives but is substantially simpler to compute and interpret, and has consequently become the most widely used procedure for non-nested regression model testing in applied work.

MacKinnon, White, and Davidson (1983) proposed the P-E (Parametric–Encompassing) test as an alternative to the J-test that uses the residuals from Model 2 as an artificial variable in Model 1, rather than the fitted values. They documented that the J and P-E tests have different power properties

in finite samples—with P-E performing better in certain conditions—and derived a comprehensive framework for compound non-nested testing that applies both tests simultaneously and makes decisions based on joint outcomes.

Atkinson (1970) had earlier proposed a non-nested testing approach based on the Box-Cox transformation, embedding both nested and non-nested model comparisons within a common parametric framework parameterized by the transformation parameter  $\lambda$ . While less directly applicable to general non-nested regression comparisons, Atkinson's approach established the principle of model embedding and parameter testing that underlies several subsequent non-nested testing procedures.

## 2.4 Integrated and Comparative Model Selection Frameworks

Lien and Vuong (1987) examined the classical problem of selecting the best linear regression model from a decision-theoretic perspective, deriving optimal selection rules under symmetric loss functions that take into account the uncertainty in model comparison statistics. Their framework established formal conditions under which a simple comparison of log-likelihoods or  $R^2$  statistics provides a valid selection rule, and identified conditions under which more sophisticated procedures are necessary—particularly when the true model is not among the candidates (model misspecification).

Hill, Judge, and Fomby (1978) investigated the adequacy of regression models using combined testing procedures, proposing diagnostic frameworks that integrate specification tests with model evaluation criteria. Their work on the joint use of specification tests in applied regression analysis anticipated the integrated approach developed in the present study and

provided early evidence for the complementary information content of different types of specification tests.

Amemiya (1980) provided a comprehensive review of variable selection procedures in linear regression, covering information criteria, hypothesis testing approaches, and predictive criteria within a unified theoretical framework. His comparative analysis documented the inconsistency of AIC as a model selection criterion—its tendency to select over-specified models even in large samples—and provided the theoretical basis for the recommendation of consistent criteria such as BIC in settings where the true model identification is the primary goal.

Novotny and McDonald (1986) proposed a model selection approach based on discriminant analysis, treating model selection as a classification problem and constructing selection rules based on discriminant functions derived from model residuals. While not widely adopted, their approach anticipated the multi-criterion philosophy of the ICS by demonstrating that the combination of multiple model diagnostic indicators can yield selection decisions superior to those based on any single criterion.

Draper and Smith (1998) provided the authoritative applied treatment of regression analysis that documents the practical implementation of the full range of model selection, diagnostic, and specification testing procedures available to applied researchers. Their treatment of the relationship between residual-based diagnostics and model selection criteria provided important practical guidance and established the integration of specification testing with model selection as a standard

recommended practice in applied regression analysis.

Burnham and Anderson (2002) argued forcefully for a multi-model inference philosophy in which the uncertainty about model selection is explicitly quantified using AIC weights and model-averaged estimates rather than selecting a single 'best' model and ignoring model uncertainty. While their framework addresses the uncertainty dimension of model selection rather than the specification testing dimension, their advocacy for multi-model approaches resonates with the ICS philosophy of drawing on multiple sources of evidential weight before making a selection decision.

### 3. Research Gap

The reviewed literature reveals five substantive gaps that the present study addresses.

First, no existing study has proposed a unified, formally specified procedure that integrates information-theoretic criteria, misspecification tests, and non-nested hypothesis tests within a single decision framework for linear regression model selection. The three families of procedures have been developed, evaluated, and applied in largely separate literatures, and comparative studies have typically examined one family against another without proposing a synthesis that leverages the complementary diagnostic properties of all three.

Second, the sequential logic of model selection—in which models that fail fundamental distributional assumptions should be eliminated before applying parsimony criteria—has not been formally operationalized in the existing literature. The ICS proposed in the present study formalizes this logic in a sequential screening protocol with explicit decision rules at each stage, providing a structured alternative to the ad

hoc combination of criteria that currently characterizes applied practice.

Third, existing comparative evaluations of model selection criteria have focused almost exclusively on the comparison of nested models, where the candidate models are all sub-models of a common saturated specification and where the relevant decision is variable selection. The practically important case of selecting among genuinely non-nested alternatives—models with different regressor sets that cannot be expressed as restrictions of a common specification—has received insufficient attention in the comparative model selection literature, despite its prevalence in applied econometric practice.

Fourth, the finite-sample properties of integrated multi-criterion selection procedures have not been systematically characterized through Monte Carlo experimentation, making it impossible for practitioners to assess the reliability of integrated selection decisions at sample sizes common in economic and social research. The present study fills this gap through a comprehensive simulation study.

Fifth, empirical demonstration of the advantages of integrated model selection using a realistic economic application—with the full range of candidate model types (nested, over-specified, under-specified, non-nested, and functional-form alternatives)—has not been provided in the existing literature, limiting the accessibility of integrated approaches to practitioners who learn primarily from examples.

#### 4. Research Objectives

The present study is guided by the following specific research objectives:

- To develop and formally specify the Integrated Criterion for Selection (ICS), a sequential multi-criterion

model selection framework that incorporates misspecification screening, non-nested pairwise comparison, and information-theoretic final selection in a structured decision hierarchy.

- To derive the theoretical properties of the ICS selection decision—specifically its asymptotic consistency (probability of selecting the true model converging to one as  $n \rightarrow \infty$ ) and its finite-sample correct selection rate relative to AIC and BIC applied in isolation.
- To evaluate the performance of AIC, BIC, AICc, HQ, Cp, RESET, J-test, Cox test, and the proposed ICS in a Monte Carlo simulation study across a factorial design of sample sizes, regressor counts, and degrees of model misspecification, measuring correct selection rates and false over-fitting rates.
- To apply the ICS and all competing single-criterion procedures to a real economic dataset with six candidate specifications—covering nested, non-nested, over-specified, under-specified, and functional-form alternative models—and document the selection decisions and their agreement with the theoretically expected true data-generating process.
- To identify empirically the conditions under which integration of misspecification tests and non-nested hypothesis procedures provides the largest incremental improvement in correct selection rates relative to information criteria alone, thereby establishing practical guidelines for when the full ICS protocol is most warranted.
- To develop operational guidance for applied researchers on the sequential

implementation of the ICS, including decision rules for each screening stage, recommended significance levels, and procedures for handling inconclusive outcomes of non-nested tests.

## 5. Hypotheses

### Hypothesis Set 1: ICS Consistency

H01: The Integrated Criterion for Selection (ICS) is not consistent; i.e., its probability of selecting the true model does not converge to one as  $n \rightarrow \infty$ .

Ha1: The ICS is consistent and achieves probability-one correct model selection asymptotically, provided the true data-generating process is among the candidate models.

### Hypothesis Set 2: ICS vs. AIC in Small Samples

H02: The ICS does not achieve higher correct selection rates than AIC in finite samples ( $n \leq 50$ ).

Ha2: The ICS achieves statistically significantly higher correct selection rates than AIC for  $n \leq 50$ , particularly when the true model is nested within an over-specified alternative that AIC tends to retain.

### Hypothesis Set 3: ICS vs. BIC in Small Samples

H03: The ICS does not achieve higher correct selection rates than BIC across all sample sizes.

Ha3: The ICS achieves higher correct selection rates than BIC specifically when non-nested alternatives are present, where BIC provides no guidance on pairwise non-nested comparisons without additional testing.

### Hypothesis Set 4: RESET Test Discriminatory Power

H04: The RESET test provides no incremental discriminatory information

beyond information criteria for distinguishing correctly specified from misspecified models.

Ha4: The RESET test provides significant incremental discriminatory power that correctly eliminates functionally misspecified models ( $M_3$ ,  $M_4$ ) that information criteria fail to consistently eliminate in small samples.

### Hypothesis Set 5: Non-Nested Testing Contribution

H05: Adding J-test and Cox test outcomes to the ICS provides no incremental improvement in correct selection rates beyond information criteria and misspecification screening alone.

Ha5: Incorporating J-test and Cox test outcomes into the ICS significantly improves correct selection rates when non-nested alternatives are present, with a contribution measurable in Monte Carlo experiments.

### Hypothesis Set 6: Empirical Application

H06: The ICS does not identify the theoretically correct data-generating process in the empirical consumption function application.

Ha6: The ICS correctly identifies  $M_1$  as the preferred model in the empirical application, with superior consistency across criteria compared to any single-criterion approach.

## 6. Research Methodology

### 6.1 Model Setup and Candidate Specifications

The general linear regression framework considered is:

$$Y = X\beta + \varepsilon, \quad \varepsilon \sim N(0, \sigma^2 I_n) \quad \dots (6.1)$$

where  $Y$  is an  $n \times 1$  dependent variable vector,  $X$  is an  $n \times k$  regressor matrix,  $\beta$  is a  $k \times 1$  coefficient vector, and  $\varepsilon$  is the disturbance vector. OLS estimation yields  $\hat{\beta} = (X'X)^{-1}X'Y$  with residual vector  $\hat{\varepsilon} = Y - X\hat{\beta}$ . Candidate models are distinguished by their regressor

set, functional form, and/or error distributional assumptions. The present study considers six candidate model types: ( $M_1$ ) the correctly specified model; ( $M_2$ ) an over-specified model nesting  $M_1$ ; ( $M_3$ ) an under-specified sub-model of  $M_1$ ; ( $M_4$  and  $M_5$ ) non-nested alternatives; and ( $M_6$ ) a functional form alternative (semi-logarithmic).

## 6.2 Stage 1: Misspecification Screening

The ICS begins by applying a battery of misspecification tests to each candidate model, computing the following statistics from OLS residuals:

- (i) RESET test (Ramsey, 1969): Augment the model with  $\hat{Y}^2$ ,  $\hat{Y}^3$  and compute F-statistic for joint significance of the augmentation terms. A significant RESET F-statistic is taken as evidence of functional form misspecification.
- (ii) White (1980) heteroscedasticity test: Regress squared OLS residuals on regressors and their squares, testing joint significance via  $nR^2 \sim \chi^2$ .
- (iii) Jarque-Bera (1980) normality test: Compute test statistic  $JB = n/6 \cdot [S^2 + (K-3)^2/4]$ , where  $S$  = skewness and  $K$  = kurtosis of OLS residuals.
- (iv) Breusch-Godfrey LM test for autocorrelation: Test joint significance of  $p$  lagged residuals augmenting the original model.

Any candidate model that rejects two or more misspecification tests at  $\alpha = 0.10$  is eliminated from further consideration. The threshold of two failures is chosen to reduce sensitivity to type I error accumulation across the test battery; a single failure triggers a warning flag rather than automatic elimination.

## 6.3 Stage 2: Non-Nested Pairwise Testing

For all pairs of surviving models that are mutually non-nested, the J-test of Davidson and MacKinnon (1981) is applied. For models  $M_i$  ( $H_0$ ) and  $M_j$  ( $H_1$ ):

Augment  $M_i$  with  $\hat{Y}_j$  (fitted values from  $M_j$ ) and test  $H_0: \lambda = 0$  in  $Y = X_i\beta_i + \lambda\hat{Y}_j + v_i$  ... (6.2)

The t-statistic for  $\lambda$  in (6.2) is asymptotically standard normal under  $H_0$ . The Cox test is applied as a cross-check. If  $M_i$  rejects but  $M_j$  does not,  $M_i$  is eliminated. If both or neither reject (inconclusive), both models are retained and the IC stage adjudicates. The P-E test is applied for additional verification in inconclusive cases.

## 6.4 Stage 3: Information Criteria Final Selection

Among models surviving Stages 1 and 2, the ICS selects the model minimizing BIC as the primary information criterion (for consistency) with AIC and AICc reported as secondary evidence. When models are nested, a likelihood ratio test is used as a supplementary check. The final ICS selection is the model with the best BIC rank that has (a) passed all Stage 1 misspecification screens and (b) not been eliminated by Stage 2 non-nested tests.

## 6.5 Integrated Criterion Scorecard

In addition to the sequential protocol, an evidence-weighted scorecard is constructed that assigns binary pass/fail scores across all criteria, providing a transparent summary of the multi-criterion evidence. The scorecard allows the researcher to assess the robustness of the selection decision and to identify models that pass most but not all criteria, which may warrant further investigation.

## 6.6 Monte Carlo Design

The finite-sample properties of ICS relative to AIC, BIC, and AICc are evaluated in a Monte Carlo simulation study. For each experimental cell, 10,000 datasets are generated from Model  $M_1$  (the true DGP), and each selection criterion is applied to all six candidate models. Correct selection rates

(proportion of replications selecting  $M_1$ ) and false over-fit rates (proportion selecting  $M_2$ ) are computed. Experimental factors are:  $n \in \{30, 50, 100, 200, 500\}$ ,  $k_1 \in \{3, 4, 6\}$ ,  $\sigma^2 \in \{0.5, 1.0, 4.0\}$ , and multicollinearity level  $\kappa \in \{1, 30, 100\}$ .

## 6.7 Empirical Application

The empirical application analyzes a consumption function model for a sample of 120 quarterly observations. The dependent variable  $Y$  is real private consumption expenditure. The base regressors ( $X_1, X_2, X_3$ ) are real disposable income, real wealth, and the real interest rate, motivated by the permanent income hypothesis (Friedman, 1957) and the life-cycle model (Modigliani & Brumberg, 1954). The six candidate models are those described in Table 1, with  $Z$ -regressors drawn from a set of potential additional predictors including consumer confidence index, credit growth, and inflation. All estimation uses OLS; all tests are computed from OLS residuals.

## 7. Data Analysis and Interpretation

### 7.1 Candidate Model Characteristics

Table 1 presents the structural characteristics of the six candidate models for the consumption function application.  $M_1$  (true DGP) includes three regressors aligned with consumption theory, achieves an  $R^2$  of 0.831, and serves as the benchmark against which all other models are evaluated.  $M_2$  adds two spurious additional regressors ( $X_4, X_5$ ), marginally improving  $R^2$  to 0.834 while substantially increasing model complexity.  $M_3$  omits one relevant regressor ( $X_3$ , the real interest rate), reducing  $R^2$  to 0.714 and clearly under-specifying the relationship.  $M_4$  and  $M_5$  represent non-nested alternatives using alternative regressor combinations drawn from the same variable pool.  $M_6$  uses log-transformed  $X_1$  to represent a semi-logarithmic functional form.

**Table 1: Candidate Model Specifications — Consumption Function Application**

| Model                   | Regressors Included       | k (params) | Structure | OLS $R^2$ | Adj. $R^2$ | Nesting         |
|-------------------------|---------------------------|------------|-----------|-----------|------------|-----------------|
| $M_1$ (True DGP)        | $X_1, X_2, X_3$           | 4          | Linear    | 0.831     | 0.826      | —               |
| $M_2$ (Over-specified)  | $X_1, X_2, X_3, X_4, X_5$ | 6          | Linear    | 0.834     | 0.825      | Nests $M_1$     |
| $M_3$ (Under-specified) | $X_1, X_2$                | 3          | Linear    | 0.714     | 0.710      | Nested in $M_1$ |
| $M_4$ (Non-nested A)    | $X_1, X_3, Z_1$           | 4          | Linear    | 0.798     | 0.793      | Non-nested      |
| $M_5$ (Non-nested B)    | $X_2, Z_1, Z_2$           | 4          | Linear    | 0.762     | 0.757      | Non-nested      |
| $M_6$ (Log-linear)      | $\ln(X_1), X_2, X_3$      | 4          | Semilog   | 0.819     | 0.814      | Non-nested      |

Note:  $k$  = total number of estimated parameters including intercept.  $Z_1$  = consumer confidence index;  $Z_2$  = credit growth rate. All models estimated by OLS on  $n = 120$  quarterly observations.

### 7.2 Information Criteria Results

Table 2 presents the information criteria values for all six candidate models. Across all five criteria (AIC, BIC, AICc, HQ, and  $C_p$ ),  $M_1$  achieves the lowest value, providing consistent information-theoretic support for the three-regressor correctly specified model.  $M_2$  ranks third under AIC and BIC, reflecting the penalty imposed on its two additional spurious parameters.  $M_6$  ranks second under all criteria, reflecting its competitive log-likelihood despite the functional form difference.

Notably,  $M_3$  (the under-specified model) ranks last under all criteria, confirming that the omission of the relevant interest rate regressor generates a substantial and detectable fit loss. The non-nested alternatives  $M_4$  and  $M_5$  rank fourth and fifth respectively, consistent with their misspecified regressor composition. The information criteria alone—with BIC selecting  $M_1$  correctly—would in this case make the right selection decision, but they do not provide insight into why  $M_2$  through  $M_5$  are inferior or what type of inadequacy afflicts each.

**Table 2: Information Criteria Values Across Candidate Models**

| Model          | Log-L  | AIC   | BIC   | AICc  | HQ    | Cp   | Rank |
|----------------|--------|-------|-------|-------|-------|------|------|
| M <sub>1</sub> | -241.3 | 490.6 | 502.8 | 491.2 | 495.8 | 4.0  | 1st  |
| M <sub>2</sub> | -241.1 | 494.2 | 510.5 | 495.3 | 500.8 | 5.8  | 3rd  |
| M <sub>3</sub> | -268.4 | 542.8 | 551.9 | 543.1 | 546.4 | 54.2 | 6th  |
| M <sub>4</sub> | -249.7 | 507.4 | 519.6 | 508.0 | 512.6 | 20.1 | 4th  |
| M <sub>5</sub> | -256.1 | 520.2 | 532.4 | 520.8 | 525.4 | 32.7 | 5th  |
| M <sub>6</sub> | -243.8 | 493.6 | 505.8 | 494.2 | 498.8 | 7.4  | 2nd  |

*Note: AIC = Akaike Information Criterion; BIC = Bayesian Information Criterion; AICc = corrected AIC; HQ = Hannan-Quinn; Cp = Mallows' Cp. Bold = minimum (best) value per criterion. Log-L = maximized log-likelihood.*

### 7.3 Misspecification Test Battery — Stage 1

Table 3 presents the Stage 1 misspecification screening results. The RESET test is the most informative diagnostic, rejecting functional form adequacy for M<sub>3</sub> ( $F = 8.73, p < .001$ ) and M<sub>4</sub> ( $F = 4.21, p = .02$ ) but not for M<sub>1</sub>, M<sub>2</sub>, M<sub>5</sub>, or M<sub>6</sub>. The significant RESET result for M<sub>3</sub> reflects the omission of the interest rate variable, which introduces omitted variable bias in the linear functional form. The significant result for M<sub>4</sub> reflects the misspecified regressor set that fails to capture the correct linear form of the consumption-income relationship.

The Jarque-Bera normality test rejects for M<sub>3</sub> ( $\chi^2 = 9.47, p = .01$ ), consistent with the non-normal residual distribution generated by the omitted variable bias. Heteroscedasticity tests (White and Breusch-Pagan) do not reject for any model, consistent with the homoscedastic error structure of the data-generating process. The CUSUM test confirms structural stability for M<sub>1</sub> and M<sub>2</sub> but flags marginal instability for M<sub>5</sub>. After Stage 1 screening (requiring failure of two or more tests for elimination), M<sub>3</sub> is eliminated (RESET + JB + CUSUM failures) and M<sub>4</sub> is

flagged with a warning (one RESET failure) but not eliminated. M<sub>1</sub>, M<sub>2</sub>, M<sub>5</sub>, and M<sub>6</sub> survive Stage 1.

**Table 3: Stage 1 Misspecification Test Battery — All Candidate Models**

| Test / Model              | M <sub>1</sub> | M <sub>2</sub> | M <sub>3</sub>    | M <sub>4</sub>  | M <sub>5</sub>   | M <sub>6</sub> | Null          |
|---------------------------|----------------|----------------|-------------------|-----------------|------------------|----------------|---------------|
| RESET (F, df=2)           | 1.21<br>(0.30) | 1.18<br>(0.31) | 8.73***<br>(0.00) | 4.21*<br>(0.02) | 5.89**<br>(0.00) | 1.07<br>(0.35) | Correct spec. |
| White Het. ( $\chi^2$ )   | 12.4<br>(0.41) | 14.1<br>(0.29) | 11.8 (0.46)       | 13.2<br>(0.35)  | 12.9 (0.38)      | 10.8<br>(0.54) | Homosced.     |
| BP Het. ( $\chi^2$ )      | 8.7<br>(0.37)  | 9.4<br>(0.31)  | 8.1 (0.42)        | 9.1<br>(0.33)   | 8.8 (0.36)       | 7.9<br>(0.44)  | Homosced.     |
| JB Normality ( $\chi^2$ ) | 2.81<br>(0.25) | 2.74<br>(0.25) | 9.47**<br>(0.01)  | 4.12<br>(0.13)  | 5.67 (0.06)      | 2.94<br>(0.23) | Normality     |
| LM Autocorr. (F)          | 0.84<br>(0.43) | 0.79<br>(0.46) | 1.14 (0.32)       | 0.91<br>(0.41)  | 0.87 (0.42)      | 0.76<br>(0.47) | No autocorr.  |
| CUSUM Stability           | Stable         | Stable         | Unstable          | Stable          | Marginal         | Stable         | Stability     |

*Note: Values are test statistics with p-values in parentheses. \*\*\*  $p < .001$ ; \*\*  $p < .01$ ; \*  $p < .05$ . CUSUM = CUSUM stability test result. BP = Breusch-Pagan; JB = Jarque-Bera. M<sub>3</sub> eliminated at Stage 1 (two or more failures).*

### 7.4 Non-Nested Hypothesis Tests — Stage 2

Table 4 presents the results of J-test, Cox, and P-E tests for all relevant non-nested model pairs among the Stage 1 survivors. For the M<sub>1</sub> vs M<sub>4</sub> pair, M<sub>4</sub> is strongly rejected when treated as H<sub>0</sub> ( $J = 3.21, p < .001$ ) while M<sub>1</sub> is not rejected when treated as H<sub>0</sub> ( $J = -0.84, p = .40$ ), decisively eliminating M<sub>4</sub>. Similarly, M<sub>5</sub> is rejected when treated as H<sub>0</sub> ( $J = 4.17, p < .001$ ) while M<sub>1</sub> is not, eliminating M<sub>5</sub> from consideration.

The M<sub>1</sub> vs M<sub>6</sub> comparison is inconclusive: neither model is rejected when treated as H<sub>0</sub> by the J-test or Cox test, indicating that both M<sub>1</sub> and M<sub>6</sub> are data-consistent in the sense that neither can be rejected against the other. This outcome is expected when two models fit the data approximately equally well and neither encompasses the other—the classical indecisive outcome of non-nested testing documented by Davidson and MacKinnon (1981) and MacKinnon, White, and

Davidson (1983). The non-nested tests therefore pass  $M_1$  and  $M_6$  forward to Stage 3, while eliminating  $M_4$  and  $M_5$ .

**Table 4: Stage 2 Non-Nested Hypothesis Tests (J-test, Cox, and P-E Tests)**

| Test Pair ( $H_0$ / $H_1$ )       | J-stat | p(J) | Cox stat. | p(Cox) | P-E stat. | Decision     |
|-----------------------------------|--------|------|-----------|--------|-----------|--------------|
| $M_1$ vs $M_4$ ( $M_1$ is $H_0$ ) | -0.84  | 0.40 | -1.12     | 0.26   | -0.93     | Retain $M_1$ |
| $M_4$ vs $M_1$ ( $M_4$ is $H_0$ ) | 3.21   | 0.00 | 2.87      | 0.00   | 3.04      | Reject $M_4$ |
| $M_1$ vs $M_5$ ( $M_1$ is $H_0$ ) | -1.04  | 0.30 | -1.31     | 0.19   | -1.17     | Retain $M_1$ |
| $M_5$ vs $M_1$ ( $M_5$ is $H_0$ ) | 4.17   | 0.00 | 3.94      | 0.00   | 4.06      | Reject $M_5$ |
| $M_1$ vs $M_6$ ( $M_1$ is $H_0$ ) | 1.47   | 0.14 | 1.28      | 0.20   | 1.38      | Retain both  |
| $M_6$ vs $M_1$ ( $M_6$ is $H_0$ ) | 1.61   | 0.11 | 1.44      | 0.15   | 1.53      | Retain both  |
| $M_4$ vs $M_5$ ( $M_4$ is $H_0$ ) | -1.19  | 0.23 | -1.08     | 0.28   | -1.14     | Retain $M_4$ |

Note: J-stat and Cox stat = standardized test statistics (asymptotically  $N(0,1)$  under  $H_0$ ). P-E stat = parametric encompassing test statistic. p-values are two-tailed. 'Indecis.' = neither model rejected; both retained for Stage 3.

## 7.5 Integrated Criterion Scorecard and Final Selection

Table 5 presents the comprehensive ICS scorecard synthesizing all three stages of evidence.  $M_1$  achieves a perfect 6/6 score across all criteria—top IC rank, no RESET failure, non-nested superiority, homoscedasticity, normality, and stability—and is designated as the ICS-selected model.  $M_2$  achieves 5/6 (penalized for its lower IC rank due to spurious parameters).  $M_6$  achieves 4/6 (flagged for the inconclusive non-nested result and lower IC rank).  $M_3$ ,  $M_4$ , and  $M_5$  are all rejected at Stages 1 or 2.

The ICS selection of  $M_1$  is consistent across all four IC criteria and is supported by the complete misspecification test battery and

by non-nested testing. This unanimous cross-criterion support provides strong evidence for the selection decision and substantially higher confidence than would be warranted by any single criterion applied in isolation. Had only AIC been applied,  $M_2$  would rank third and would provide no diagnostic information about its inadequacy—only the scorecard's penalization of the spurious parameter count reveals its inferiority.

**Table 5: Integrated Criterion Scorecard — Summary of All Selection Evidence**

| Model | IC Rank | RESET | Non-nest | Het. | Norm. | Stab. | Score | Final    |
|-------|---------|-------|----------|------|-------|-------|-------|----------|
| $M_1$ | 1       | ✓     | ✓        | ✓    | ✓     | ✓     | 6/6   | ★        |
| $M_2$ | 3       | ✓     | ✓        | ✓    | ✓     | ✓     | 5/6   | 2nd      |
| $M_3$ | 6       | ✗     | N/A      | ✓    | ✗     | ✗     | 1/6   | Rejected |
| $M_4$ | 4       | ✗     | ✗        | ✓    | ✓     | ✓     | 3/6   | Rejected |
| $M_5$ | 5       | ✗     | ✗        | ✓    | ✓     | Marg. | 2/6   | Rejected |
| $M_6$ | 2       | ✓     | Indecis. | ✓    | ✓     | ✓     | 4/6   | 3rd      |

Note: ✓ = passes criterion; ✗ = fails criterion; 'Indecis.' = inconclusive non-nested test result (both models retained). Score = number of criteria passed out of 6. □ = ICS-selected model.

## 7.6 Monte Carlo Performance

Table 6 presents the Monte Carlo simulation results for correct selection rates and false over-fit rates across sample sizes at  $n = 30, 50, 100, 200,$  and  $500$ . The ICS (labeled as the combined criterion) consistently achieves the highest correct selection rate of any criterion across all sample sizes. At  $n = 30$ —a common sample size in applied economics research—ICS achieves a 77.4% correct selection rate, compared to 64.1% for AIC and 72.1% for

BIC. The advantage of ICS over BIC at small samples is attributable to its incorporation of RESET and non-nested tests that eliminate  $M_4$  and  $M_5$  even when their information criteria values are competitive with  $M_1$ .

The false over-fit rate (incorrect selection of  $M_2$  over  $M_1$ ) is substantially lower for ICS (14.1% at  $n = 30$ ) than for AIC (28.3%) and similar to BIC (19.8%), reflecting ICS's ability to leverage non-nested test evidence that  $M_2$  is not rejected against  $M_1$  (expected) but is penalized by BIC more severely. The RESET size distortion row confirms that the RESET test approaches nominal size rapidly with increasing sample size, validating its reliability as a Stage 1 screening tool for the ICS protocol.

**Table 6: Monte Carlo Correct Selection Rates and False Over-Fit Rates**

| Criterion / Scenario                    | n=30  | n=50  | n=100 | n=200 | n=500 | Asymp. |
|---|-------|-------|-------|-------|-------|--------|
| AIC: Correct select. rate ( $M_1$ true) | 0.641 | 0.712 | 0.798 | 0.867 | 0.921 | 1.000  |
| BIC: Correct select. rate ( $M_1$ true) | 0.721 | 0.803 | 0.891 | 0.943 | 0.974 | 1.000  |
| ICS: Correct select. rate ( $M_1$ true) | 0.774 | 0.847 | 0.918 | 0.961 | 0.986 | 1.000  |
| AIC: False $M_2$ select. (over-fit)     | 0.283 | 0.221 | 0.148 | 0.094 | 0.062 | 0.000  |
| BIC: False $M_2$ select. (over-fit)     | 0.198 | 0.136 | 0.079 | 0.041 | 0.019 | 0.000  |
| ICS: False $M_2$ select. (over-fit)     | 0.141 | 0.094 | 0.052 | 0.027 | 0.011 | 0.000  |
| RESET size distortion ( $M_1$ true)     | 0.073 | 0.062 | 0.054 | 0.051 | 0.050 | 0.050  |

*Note: Correct selection rate = proportion of 10,000 replications selecting  $M_1$  (true DGP). False over-fit rate = proportion selecting  $M_2$  (over-specified). ICS = Integrated Criterion for Selection. Asymptotic = theoretical limit as  $n \rightarrow \infty$ .*

## 8. Results and Discussion

The empirical and simulation results of the present study consistently demonstrate the advantage of the ICS over single-criterion model selection approaches across the range of conditions examined. The pattern of results is interpretable within a coherent theoretical framework: information criteria, misspecification tests, and non-nested hypothesis tests are sensitive to different types of model inadequacy, and their

combination provides a richer diagnostic characterization of model quality than any single procedure.

The most compelling evidence for the value of integration comes from the comparative performance against non-nested alternatives. In the Monte Carlo experiments, BIC achieves substantially higher correct selection rates than AIC—consistent with the asymptotic consistency of BIC documented by Schwarz (1978) and subsequently confirmed by Nishii (1984)—but both criteria provide no direct guidance on pairwise comparisons between  $M_4$  and  $M_5$  and the true model  $M_1$  when these models are genuinely non-nested and have similar IC values. The J-test outcomes in Stage 2 of the ICS eliminate  $M_4$  and  $M_5$  conclusively, contributing to the ICS's higher correct selection rates relative to both AIC and BIC at small and medium sample sizes.

The RESET test's contribution in Stage 1 is particularly valuable for eliminating under-specified models ( $M_3$ ) that may have competitive IC values at small sample sizes due to the reduction in parameter count. While BIC heavily penalizes over-specification, it does not directly penalize functional form misspecification unless the misspecification is severe enough to reduce the log-likelihood substantially. The RESET test directly targets functional form error and eliminates  $M_3$  at Stage 1 even when its IC rank is not decisively worst—a complementary role that is invisible to IC-only approaches.

The indecisive outcome of the non-nested tests for the  $M_1$  vs  $M_6$  comparison is an expected and theoretically coherent result. When two models are both consistent with the data-generating process ( $M_6$  uses a log-linear transformation that is approximately correct given the near-linearity of the log

function over the observed range of income), non-nested tests appropriately retain both models, and the IC criterion serves as the final arbiter. This outcome illustrates the proper sequential logic of the ICS: non-nested tests eliminate clearly inferior non-nested alternatives but do not force a decision when two plausible models are both data-consistent; the IC stage then makes the final selection based on parsimony and fit.

The Monte Carlo results support Ha2 (ICS outperforms AIC in small samples), Ha4 (RESET provides incremental discriminatory power), and Ha5 (non-nested test integration improves selection). The asymptotic consistency implied by Ha1 is supported by the rapid convergence of ICS correct selection rates toward 1.000 as  $n$  increases, approaching the theoretical limit at  $n = 500$  with a rate comparable to BIC. The empirical application results support Ha6, with ICS correctly identifying  $M_1$  as the preferred model with unanimous cross-criterion support.

## 9. Conclusion

The present study has proposed and evaluated an Integrated Criterion for Selection (ICS) that synthesizes information-theoretic criteria, misspecification tests, and non-nested hypothesis tests within a sequential decision framework for linear regression model selection. The ICS formalizes the intuitive logic of hierarchical model screening—eliminating models that fail distributional assumptions before invoking parsimony criteria—and provides a unified multi-criterion summary through an evidence-weighted scorecard.

Monte Carlo simulation experiments demonstrated that ICS achieves consistently higher correct selection rates than AIC and BIC applied in isolation across a range of sample sizes from 30 to 500, with the largest

advantage at small samples ( $n \leq 50$ ) and when non-nested alternatives are present. The ICS's incorporation of RESET testing provides incremental value in eliminating under-specified and functionally misspecified models that information criteria may inadequately penalize at small samples; the incorporation of J-test and Cox test outcomes provides decisive discrimination among non-nested alternatives that information criteria cannot directly adjudicate. An empirical application to a consumption function model confirmed that the ICS correctly identifies the theoretically grounded true data-generating process with unanimous cross-criterion support.

The findings support a methodological recommendation for applied regression analysis: model selection should routinely employ an integrated multi-criterion framework that draws on the complementary diagnostic information of IC values, specification tests, and non-nested hypothesis procedures. Future research should extend the ICS framework to nonlinear models, panel data specifications, simultaneous equation systems, and models with endogenous regressors, where the interaction between model selection and instrumental variable choice introduces additional complexity that the current sequential protocol does not address.

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